

What can be discovered in a formula for the variance of the BLUE

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Linear model:

$$\begin{aligned}y &= F\theta + \varepsilon; & \theta &\in \mathbb{R}^p \\ E(\varepsilon) &= 0, \text{Var}(\varepsilon) = \sigma^2 I\end{aligned}$$

the variance of the BLUE for $c^T\theta$ is

$$\text{Var}(c^T \hat{\theta}) = \sigma^2 c^T M^{-1} c = \sigma^2 \max_{\alpha: M\alpha \neq 0} \frac{(c^T \alpha)^2}{\alpha^T M \alpha}$$

where $M = F^T F$ is the information matrix
(cf. textbooks of J. Anděl or of C.R. Rao)

Results from textbooks

Design in linear models:

$$y_x = f^T(x)\theta + \varepsilon_x; \quad \mathbf{x} \in \mathcal{X} = \text{design space}$$
$$E(\varepsilon_x) = 0, \quad \text{Var}(\varepsilon_x) = \sigma^2 \equiv 1$$

x_1, \dots, x_N = exact design

relative frequency = ξ = **design**

$$\text{Var}_\xi \left(\mathbf{c}^T \hat{\theta} \right) = \max_{\alpha: M(\xi)\alpha \neq 0} \frac{(\mathbf{c}^T \alpha)^2}{\alpha^T M(\xi) \alpha}$$
$$M(\xi) = \int_{\mathcal{X}} f(x) f^T(x) d\xi(x)$$

The present talk is about **two extensions** of the c-optimality criterion

$$\xi \rightarrow \max_{\alpha: M(\xi)\alpha \neq 0} \frac{(\mathbf{c}^T \alpha)^2}{\alpha^T M(\xi) \alpha}$$

to other models.

Extension to **parameter-free** (infinite-dimensional) lin. models

\mathcal{H} is the set (linear space) of all allowed regression functions defined on the set \mathcal{X} (for example $\mathcal{H} = \{\eta(\cdot) = f^T(\cdot)\theta, \theta \in \mathbb{R}^p\}$).

Further,

$$g : \mathcal{H} \rightarrow \mathbb{R}$$

is a linear functional,

$$\|\eta\|_{\xi}^2 = \int_{\mathcal{X}} \eta^2(\mathbf{x}) d\xi(\mathbf{x}); \quad \eta \in \mathcal{H}$$

Define

$$\Phi_g(\xi) \equiv \max_{\eta \in \mathcal{H}} \frac{g^2(\eta)}{\|\eta\|_{\xi}^2}.$$

Remark: THERE IS NO INFORMATION MATRIX IN $\Phi_g(\xi)$.

Extension for the **nonlinearly** parametrized regression models

$$\begin{aligned}y_x &= \eta(\mathbf{x}, \theta) + \varepsilon_x; \quad \mathbf{x} \in \mathcal{X} \\ E(\varepsilon_x) &= 0, \quad \text{Var}(\varepsilon_x) = \sigma^2 \equiv 1\end{aligned}$$

$h(\theta)$ = (smooth) function to be estimated

$\bar{\theta}$ = the true (guessed) value of θ

Define

$$\begin{aligned}\Psi_h(\xi) &= \min_{\theta \in \mathbb{R}^p, h(\theta) \neq h(\bar{\theta})} \psi_h(\xi; \theta) \\ \text{with } \psi_h(\xi; \theta) &= \frac{\int_{\mathcal{X}} [\eta(\mathbf{x}, \theta) - \eta(\mathbf{x}, \bar{\theta})]^2 d\xi(\mathbf{x})}{[h(\theta) - h(\bar{\theta})]^2}\end{aligned}$$

Again: THERE IS NO INFORMATION MATRIX IN $\Psi_h(\xi)$.

Property of Extension 1:

If the model is linear and parametrized, and if

$$g \left[f^T (\cdot) \theta \right] = c^T \theta$$

then

$$\Phi_g(\xi) = \text{Var}_\xi \left(c^T \hat{\theta} \right).$$

Property of Extension 2:

If the model is linear and parametrized, and if

$$\begin{aligned} \det [M(\xi)] &\neq 0, \\ h(\theta) &= c^T \theta + d \end{aligned}$$

then

$$\Psi_h(\xi) = \left[\text{Var}_\xi \left(c^T \hat{\theta} \right) \right]^{-1}.$$

Basic properties

Proof:

When

$$\begin{aligned}\eta(\mathbf{x}, \alpha) &= \mathbf{f}^T(\mathbf{x})\alpha, \\ \mathbf{g}[\eta(\cdot, \alpha)] &= \mathbf{c}^T\alpha,\end{aligned}$$

then

$$\begin{aligned}\alpha^T M(\xi)\alpha &= \alpha^T \left[\int_{\mathcal{X}} \mathbf{f}(\mathbf{x}) \mathbf{f}^T(\mathbf{x}) d\xi(\mathbf{x}) \right] \alpha \\ &= \int_{\mathcal{X}} \left[\mathbf{f}^T(\mathbf{x})\alpha \right]^2 d\xi(\mathbf{x}) = \|\eta(\cdot, \alpha)\|_{\xi}^2.\end{aligned}$$

Hence

$$\max_{\eta \in \mathcal{H}} \frac{\mathbf{g}^2(\eta)}{\|\eta\|_{\xi}^2} = \max_{\alpha: M(\xi)\alpha \neq 0} \frac{(\mathbf{c}^T\alpha)^2}{\alpha^T M(\xi)\alpha}.$$

Basic properties

Similarly, when

$$\begin{aligned}h(\theta) &= \mathbf{c}^T \theta + \mathbf{d}, \\ \eta(\mathbf{x}, \theta) - \eta(\mathbf{x}, \bar{\theta}) &= \mathbf{f}^T(\mathbf{x}) (\theta - \bar{\theta}),\end{aligned}$$

then

$$\begin{aligned}& \min_{\theta \in \mathbb{R}^p, h(\theta) \neq h(\bar{\theta})} \frac{\int_{\mathcal{X}} [\eta(\mathbf{x}, \theta) - \eta(\mathbf{x}, \bar{\theta})]^2 d\xi(\mathbf{x})}{[h(\theta) - h(\bar{\theta})]^2} = \\&= \min_{\theta \in \mathbb{R}^p, \mathbf{c}^T(\theta - \bar{\theta}) \neq 0} \frac{(\theta - \bar{\theta})^T M(\xi) (\theta - \bar{\theta})}{[\mathbf{c}^T (\theta - \bar{\theta})]^2} = \\&= \left[\max_{\alpha \neq 0} \frac{(\mathbf{c}^T \alpha)^2}{\alpha^T M(\xi) \alpha} \right]^{-1}.\end{aligned}$$



The interpretation of $\Phi_g(\xi)$

Take

$x_1, \dots, x_N =$ exact design,

$\xi =$ “frequency” design.

Define the **observed random measure**

$$Y_\xi(B) \equiv \frac{1}{N} \sum_{i=1, x_i \in B}^N y_{x_i}.$$

Then for any $\eta \in \mathcal{H}$

$$\begin{aligned} E_\eta(Y_\xi(B)) &= \int_B \eta(x) d\xi(x), \\ \text{Cov}[Y_\xi(B_1), Y_\xi(B_2)] &= \xi(B_1 \cap B_2). \end{aligned}$$

A linear estimator is an integral with respect to this observed random measure.

The interpretation of $\Phi_g(\xi)$

Theorem (cf. A. Pázman, Kybernetika 1978)

The functional g is linearly estimable under ξ iff there is a function

$$I \in L_2(\xi)$$

such that

$$g(\eta) = \int_{\mathcal{X}} I(x) \eta(x) d\xi(x); \quad \eta \in \mathcal{H}.$$

Then

$$\text{Var}_{\xi}(\text{BLUE}) = \max_{\eta \in \mathcal{H}} \frac{g^2(\eta)}{\|\eta\|_{\xi}^2} = \|I_{\xi}\|_{\xi}^2,$$

where $I_{\xi} =$ the $L_2(\xi)$ -orthogonal projection of I onto \mathcal{H} .

The interpretation of $\Psi_h(\xi)$

If $\|\theta - \bar{\theta}\|$ is small, then

$$\begin{aligned}\Psi_h(\xi) &= \min_{\theta} \frac{\int_{\mathcal{X}} [\eta(\mathbf{x}, \theta) - \eta(\mathbf{x}, \bar{\theta})]^2 d\xi(\mathbf{x})}{[h(\theta) - h(\bar{\theta})]^2} \\ &= \min_{\theta} \frac{\int_{\mathcal{X}} \left[\frac{\partial \eta(\mathbf{x}, \theta)}{\partial \theta} \Big|_{\bar{\theta}} (\theta - \bar{\theta}) + o(\|\theta - \bar{\theta}\|) \right]^2 d\xi(\mathbf{x})}{\left[\frac{\partial h(\theta)}{\partial \theta} \Big|_{\bar{\theta}} (\theta - \bar{\theta}) + o(\|\theta - \bar{\theta}\|) \right]^2} \\ &= \min_{\theta} \frac{\int_{\mathcal{X}} (\theta - \bar{\theta})^T M(\xi, \bar{\theta}) (\theta - \bar{\theta}) + o(\|\theta - \bar{\theta}\|)}{\left[\frac{\partial h(\theta)}{\partial \theta} \Big|_{\bar{\theta}} (\theta - \bar{\theta}) \right]^2} \\ &= \left\{ \left[\frac{\partial h(\theta)}{\partial \theta^T} \right]_{\bar{\theta}} M^{-1}(\xi, \bar{\theta}) \left[\frac{\partial h(\theta)}{\partial \theta} \right]_{\bar{\theta}} \right\}^{-1} + o(\|\theta - \bar{\theta}\|)\end{aligned}$$

\Rightarrow standard local c-optimality criterion in nonlinear models.

The interpretation of $\Psi_h(\xi)$

However, there is much more in $\Psi_h(\xi)$, namely

$$\Psi_h(\xi) = \min_{\theta \in \mathbb{R}^p, h(\theta) \neq h(\bar{\theta})} \frac{\int_{\mathcal{X}} [\eta(\mathbf{x}, \theta) - \eta(\mathbf{x}, \bar{\theta})]^2 d\xi(\mathbf{x})}{[h(\theta) - h(\bar{\theta})]^2}$$

detects the lack of identifiability:

1) If

$$\int_{\mathcal{X}} [\eta(\mathbf{x}, \theta) - \eta(\mathbf{x}, \bar{\theta})]^2 d\xi(\mathbf{x}) = 0$$

⇓

we miss a standard condition for consistency and asymptotical normality of $\hat{\theta}$.

2)

If $\int_{\mathcal{X}} [\eta(\mathbf{x}, \theta) - \eta(\mathbf{x}, \bar{\theta})]^2 d\xi(\mathbf{x})$ much smaller than $[h(\theta) - h(\bar{\theta})]^2$

⇓

the estimator $h(\hat{\theta})$ can not distinguish $h(\bar{\theta})$ from $h(\theta)$.

Convexity, concavity and directional derivatives

Theorem

The function $\Phi_g(\xi)$ is convex in ξ .

The function $\Psi_h(\xi)$ is concave in ξ .

Proof:

The concavity of

$$\Psi_h(\xi) = \min_{\theta \in \mathbb{R}^p, h(\theta) \neq h(\bar{\theta})} \psi_h(\xi; \theta)$$

is evident, since

$$\psi_h(\xi; \theta) = \frac{\int_{\mathcal{X}} [\eta(\mathbf{x}, \theta) - \eta(\mathbf{x}, \bar{\theta})]^2 d\xi(\mathbf{x})}{[h(\theta) - h(\bar{\theta})]^2}$$

is linear in ξ .

Convexity, concavity and directional derivatives

Further, for every $\eta \in \mathcal{H}$

$$0 \leq \|l_\xi - \eta\|_\xi^2 = \|l_\xi\|_\xi^2 - 2g(\eta) + \|\eta\|_\xi^2.$$

Hence

$$\Phi_g(\xi) = \|l_\xi\|_\xi^2 \geq 2g(\eta) - \|\eta\|_\xi^2$$

and the equality is attained if $\eta = l_\xi$.

$$\Phi_g(\xi) = \max_{\eta \in \mathcal{H}} \left\{ 2g(\eta) - \int_{\mathcal{X}} \eta^2(x) d\xi(x) \right\}$$

which is evidently convex in ξ .



Convexity, concavity and directional derivatives

A convex (concave) function has always a **directional derivative** (at the point ξ and in a direction given by δ)

$$\nabla_{\delta} \Phi_g(\xi) = \lim_{\beta \rightarrow 0} \frac{\Phi_g[(1 - \beta)\xi + \beta\delta] - \Phi_g(\xi)}{\beta}$$

and similarly for $\Psi_h(\xi)$.

As well known, a design μ is Φ_g -optimal if and only if

$$\nabla_{\delta} \Phi_g(\mu) \geq 0, \quad \forall \delta$$

Similarly for the Ψ_h -optimality we require

$$\nabla_{\delta} \Psi_h(\mu) \leq 0, \quad \forall \delta$$

Convexity, concavity and directional derivatives

Theorem

$$\begin{aligned} \text{We have } \nabla_{\delta} \Phi_g(\xi) &= \int_{\mathcal{X}} l_{\xi}^2(\mathbf{x}) d\xi(\mathbf{x}) - \int_{\mathcal{X}} l_{\xi}^2(\mathbf{x}) d\delta(\mathbf{x}), \\ \nabla_{\delta} \Psi_h(\xi) &= \min_{\theta \in \Theta(\xi)} \psi_h(\delta, \theta) - \Psi_h(\xi), \\ \Theta(\xi, h) &= \arg \min_{\theta \in \Theta} \psi_h(\xi, \theta). \end{aligned}$$

The first expression requires to compute a projection in $L_2(\xi)$, the second requires to solve a minimization problem to obtain $\Theta(\xi, h)$.

Proof: From books on minimax solutions we obtain

$$\begin{aligned} \nabla_{\delta} \left[\min_{\theta \in \Theta} \psi_h(\xi, \theta) \right] &= \min_{\theta \in \Theta(\xi, h)} [\nabla_{\delta} \psi_h(\xi, \theta)] \\ &= \min_{\theta \in \Theta(\xi, h)} \frac{d}{d\beta} [\psi_h((1 - \beta)\xi(\mathbf{x}) + \beta\delta, \theta)] \\ &= \min_{\theta \in \Theta(\xi, h)} [\psi_h(\delta, \theta) - \psi_h(\xi, \theta)] \end{aligned}$$

since $\psi_h(\xi, \theta)$ is linear in ξ .

Convexity, concavity and directional derivatives

Similarly, from

$$\Phi_g(\xi) = \max_{\eta \in \mathcal{H}} \left\{ 2g(\eta) - \int_{\mathcal{X}} \eta^2(x) d\xi(x) \right\}$$

we have

$$\begin{aligned} & \nabla_{\delta} \left[\max_{\eta \in \mathcal{H}} \left\{ 2g(\eta) - \int_{\mathcal{X}} \eta^2(x) d\xi(x) \right\} \right] \\ &= \max_{\eta \in \mathcal{H}(\xi)} \left[\nabla_{\delta} \left(- \int_{\mathcal{X}} \eta^2(x) d\xi(x) \right) \right] \end{aligned}$$

where

$$\begin{aligned} \mathcal{H}(\xi) &= \arg \max_{\eta \in \mathcal{H}} \left\{ 2g(\eta) - \int_{\mathcal{X}} \eta^2(x) d\xi(x) \right\} = \{\eta_{\xi}\}, \\ \nabla_{\delta} \left(\int_{\mathcal{X}} \eta^2(x) d\xi(x) \right) &= \int_{\mathcal{X}} \eta^2(x) d\delta(x) - \int_{\mathcal{X}} \eta^2(x) d\xi(x). \end{aligned}$$

□

Theorem (Equivalence for Φ_g in parameter-free models)

A design μ is Φ_g optimal if and only if

$$\max_{x \in \mathcal{X}} I_{\mu}^2(x) = \int_{\mathcal{X}} I_{\mu}^2(x) d\mu(x)$$

Proof:

$$\nabla_{\delta} \Phi_g(\xi) = \int_{\mathcal{X}} I_{\xi}^2(x) d\xi(x) - \int_{\mathcal{X}} I_{\xi}^2(x) d\delta(x) \geq 0; \quad \forall \delta$$

□

Particular case: linear model with $\det M(\mu) \neq 0$

$$I_{\mu}(x) = c^T M^{-1}(\mu) f(x)$$
$$\Downarrow$$

$$\max_{x \in \mathcal{X}} \left[c^T M^{-1}(\mu) f(x) \right]^2 = c^T M^{-1}(\mu) c$$

Theorem (Equivalence for Ψ_h in nonlinear models)

Suppose that $\Theta(\mu, h) = \{\theta_{\mu, h}\} = \text{a unique point}$.

Then a design μ is Ψ_h -optimal if and only if

$$\max_{x \in \mathcal{X}} [\eta(x, \theta_{\mu, h}) - \eta(x, \bar{\theta})]^2 = \int_{\mathcal{X}} [\eta(x, \theta_{\mu, h}) - \eta(x, \bar{\theta})]^2 d\mu(x)$$

Proof:

$$\forall \delta \quad \nabla_{\delta} \Psi_h(\mu) = \psi_h(\delta, \theta_{\mu, h}) - \psi_h(\xi, \theta_{\mu, h}) \leq 0$$

\Downarrow

$$\forall \delta \quad \frac{\int_{\mathcal{X}} [\eta(x, \theta_{\mu, h}) - \eta(x, \bar{\theta})]^2 d\delta(x)}{[h(\theta_{\mu, h}) - h(\bar{\theta})]^2} \leq$$

$$\leq \Psi_h(\mu) = \frac{\int_{\mathcal{X}} [\eta(x, \theta_{\mu, h}) - \eta(x, \bar{\theta})]^2 d\mu(x)}{[h(\theta_{\mu, h}) - h(\bar{\theta})]^2}$$

□

Particular case: linear model with $\det M(\mu) \neq 0$

$$\begin{aligned}\theta_{\mu,h} - \bar{\theta} &= M^{-1}(\mu) \mathbf{c} \\ \eta(\mathbf{x}, \theta_{\mu,h}) - \eta(\mathbf{x}, \bar{\theta}) &= \mathbf{f}^T(\mathbf{x}) M^{-1}(\mu) \mathbf{c} \\ &\Downarrow \\ \max_{\mathbf{x} \in \mathcal{X}} \left[\mathbf{f}^T(\mathbf{x}) M^{-1}(\mu) \mathbf{c} \right]^2 &= \int_{\mathcal{X}} \left[\mathbf{f}^T(\mathbf{x}) M^{-1}(\mu) \mathbf{c} \right]^2 d\mu(\mathbf{x}) \\ &= \mathbf{c}^T M^{-1}(\mu) \mathbf{c}\end{aligned}$$

Remark: If $\Theta(\mu, h)$ has finitely many points we have:

Theorem

A design μ is Ψ_h optimal if and only if there is a probability measure λ on $\Theta(\mu, h)$ such that $\sum_{\theta \in \Theta(\mu, h)} \lambda(\theta) [\psi_h(\delta, \theta) - \psi_h(\mu, \theta)] \leq 0$; $\forall \delta$, or, equivalently,

$$\max_{\mathbf{x} \in \mathcal{X}} \sum_{\theta \in \Theta(\mu, h)} \lambda(\theta) [\eta(\mathbf{x}, \theta) - \eta(\mathbf{x}, \bar{\theta})]^2 = \Psi_h(\mu).$$

Extending Elfving's theorem

Theorem (Elfving's theorem (1952) in parametrized linear models)

In the linear model $\eta(x, \theta) = f^T(x)\theta$ let

$$S \equiv \text{co} \{f(\mathcal{X}) \cup [-f(\mathcal{X})]\}.$$

Then

$$\mu \in \arg \min_{\xi} \text{Var}_{\xi} \left(c^T \hat{\theta} \right)$$

if and only if $\exists \beta \in \mathbb{R}$ such that

- i) $\beta c \in \text{boundary of } S$
- ii) $\beta c = \int_{\mathcal{X}} \delta(x) f(x) d\mu(x)$ for some $\delta(x) \in \{-1, 1\}$

Extensions to other criteria by Studden, Dette etc.

Numerical computation:

- J. López Fidalgo & Rodríguez-Díaz (Metrika)
- R. Harman & T. Jurík (Comp. Stat. & Data Analysis, 2008): a linear programming method

Extension of Elfving's theorem to Φ_g in parameter-free models

Theorem (cf. A. Pázman, Kybernetika 1978)

Hahn-Banach theorem \Rightarrow the functional g can be extended from \mathcal{H} to the set $C(\mathcal{X})$ so that

$$\max_{\eta \in \mathcal{H}} \frac{[g(\eta)]^2}{\|\eta\|_\infty^2} = \max_{f \in C(\mathcal{X})} \frac{[g(f)]^2}{\|f\|_\infty^2}.$$

Riesz theorem \Rightarrow there is a generalized measure ν on \mathcal{X} such that

$$g(f) = \int_{\mathcal{X}} f(x) d\nu(x) \quad ; f \in C(\mathcal{X}).$$

Let $\nu = \nu^+ - \nu^-$ be the Jordan decomposition of ν .

Then the design $\mu = \frac{\nu^+ + \nu^-}{\nu^+(\mathcal{X}) + \nu^-(\mathcal{X})}$ is Φ_g -optimal.

Some words from G.B. Shaw before the today's evening to which you are all cordially invited

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Imagination is the beginning of creation. You imagine what you desire, you will what you imagine and at last you create what you will.